

9.1 1/3 CH 9 HYPOTHESIS TESTING

9.1

A HYPOTHESIS TEST IS USED TO MAKE A DECISION ABOUT A POPULATION PARAMETER.

[HANDOUT 10] #1. "WANT" $\mu = 1.00$ IN, $\mu \neq 1.00$ IS "UNDESIRABLE"

NOTE

$H_0: \mu = 1.00$ NULL HYPOTHESIS

$H_a: \mu \neq 1.00$ ALTERNATIVE HYPOTHESIS

WE TEST H_0 VS. H_a .

THE RESULT WILL SUGGEST SOME CAUSE OF ACTION.

IN OUR EXAMPLE, IF WE REJECT H_0 IN FAVOR OF H_a , WE WILL HAVE TO FIX THE MACHINE.

TO PERFORM TEST: WE TAKE A RANDOM SAMPLE AND SEE WHICH OF THE TWO HYPOTHESES OUR DATA IS MOST CONSISTENT WITH.

THE STRATEGY IS "INNOCENT UNTIL PROVEN GUILTY" FOR H_0 .

IF "GUILTY", WE "REJECT H_0 IN FAVOR OF H_a ".

IF "NOT GUILTY", WE "FAIL TO REJECT H_0 ".

#1. $n = 50$ BOOTS, $\bar{X} = 1.02$ ", $S = 0.04$ "

IS $\bar{X} = 1.02$ CLOSE ENOUGH TO $\mu = 1$?

IF THE MEAN LENGTH OF POPULATION IS TRULY $\mu = 1$,

THEN $\bar{X} \sim N(1.00, \frac{\sigma^2}{n}) = N(1, \frac{.04^2}{50})$.

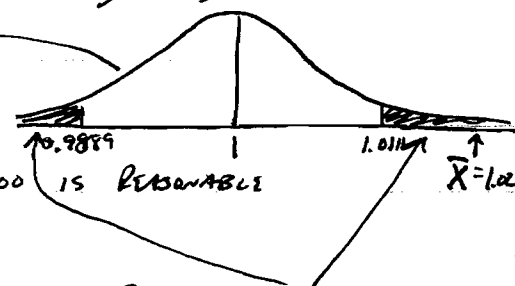
SET UP A $100(1-\alpha)\%$ CI ABOUT $\mu = 1.00$.

IF OBSERVED \bar{X} FALLS IN THE INTERVAL, $\mu = 1.00$

SO WE "FAIL TO REJECT H_0 ".

IF OBSERVED \bar{X} FALLS OUTSIDE THE INTERVAL IN THE REJECTION REGION,

WE "REJECT H_0 IN FAVOR OF H_a ".

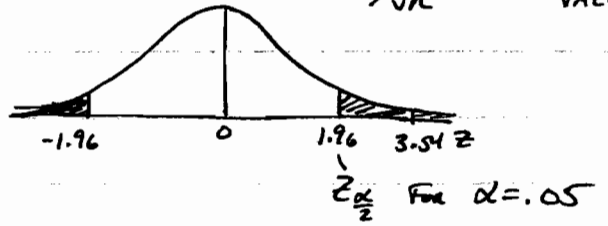


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STANDARDIZE TO $N(0, 1)$

Z-SCORE IS $Z = \frac{\bar{X} - \mu_0}{S/\sqrt{n}}$, THE STANDARDIZED OBSERVED VALUE.

$$Z = \frac{1.02 - 1}{.04/\sqrt{50}} = 3.54$$



SINCE Z IS IN THE REJECTION REGION,
WE REJECT H_0 IN FAVOR OF H_a ,
CONCLUDING THAT $\mu \neq 1$.

POSSIBLE HYPOTHESES

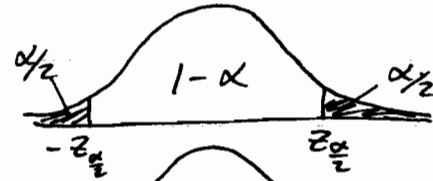
TWO-SIDED $H_0: \mu = \mu_0$
 $H_a: \mu \neq \mu_0$

ONE-SIDED $H_0: \mu = \mu_0$
 $H_a: \mu < \mu_0$

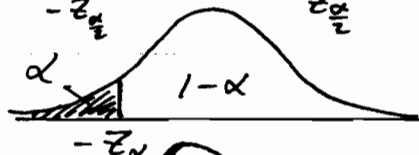
ONE-SIDED $H_0: \mu = \mu_0$
 $H_a: \mu > \mu_0$

REJECTION REGION

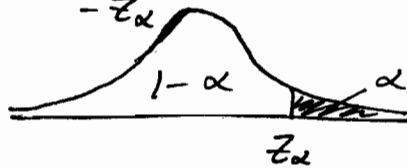
REJECT IF



$Z < -z_{\alpha/2}$ OR
 $Z > z_{\alpha/2}$



$Z < -z_\alpha$



$Z > z_\alpha$

TYPICAL VALUES FOR α , THE SIGNIFICANCE LEVEL OF THE TEST,

$\alpha = .10$

$= .05$

$= .01$

WE CAN MAKE TWO TYPES OF ERRORS

TYPE I ERROR, REJECT H_0 WHEN H_0 IS TRUE

TYPE II ERROR, FAIL TO REJECT H_0 WHEN H_0 IS FALSE.

DECISION	STATE OF NATURE	
	H_0 TRUE	H_0 FALSE
FAIL TO REJECT H_0	CORRECT DECISION	TYPE II ERROR
REJECT H_0	TYPE I ERROR	CORRECT DECISION

TYPICAL FIX PROBABILITY OF MAKING A TYPE I ERROR = α .

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IN EXAMPLE 1, WE REJECTED H_0 AT THE $\alpha = .05$ LEVEL WITH $Z_{\alpha/2} = 1.96$.
WE WERE WILLING TO ACCEPT A 5% CHANCE OF REJECTING H_0 ~~WHEN~~ ^{IF}
 H_0 IS TRUE.

THE PROBABILITY OF MAKING A TYPE II ERROR = β ,

THE PROBABILITY OF FAILING TO REJECT H_0 WHEN H_0 IS FALSE.

THE POWER OF THE TEST = $1 - \beta$

THE PROBABILITY OF REJECTING H_0 WHEN IT IS FALSE.

THE POWER IS COMPUTED FOR SPECIFIC VALUES OF THE PARAMETER
IN THE ALTERNATIVE HYPOTHESIS.

EX FIND THE POWER OF THE HYPOTHESIS TEST WHEN THE TRUE VALUE OF $\mu = 1.01$ IN.

$\beta = P_r(\text{DO NOT REJECT } H_0: \mu = 1.00 \text{ WHEN } H_0 \text{ IS FALSE } (\mu = 1.01))$

LET $\alpha = 0.05$.

ASSUME H_0 IS TRUE, THEN $\bar{X} \sim N(1.00, \frac{\sigma^2}{n})$.

A 95% CI IS $(.9889 \leq \bar{X} \leq 1.0111)$.

WE DO NOT REJECT WHEN \bar{X} IS IN THIS INTERVAL.

WE WANT $\beta = P_r(\text{DO NOT REJECT} \mid H_0 \text{ IS FALSE } (\mu = 1.01))$
 $= P_r(.9889 \leq \bar{X} \leq 1.0111 \mid \mu = 1.01)$ ^{SHIFT}

IT IS THE AREA UNDER $N(1.01, \frac{\sigma^2}{n})$

INSIDE THE 95% CI OF $N(1, \frac{\sigma^2}{n})$.

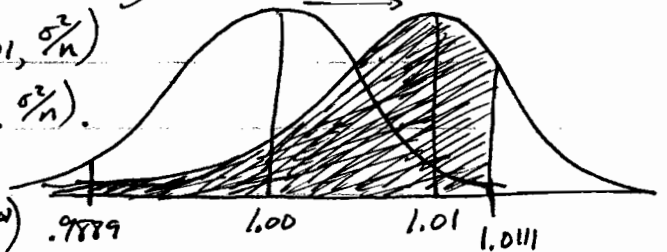
$$P_r\left(\frac{.9889 - 1.01}{\frac{\sigma}{\sqrt{n}}} \leq \frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \leq \frac{1.0111 - 1.01}{\frac{\sigma}{\sqrt{n}}}\right)$$

$$= P_r(-3.73 \leq Z \leq 0.194)$$

$$= .575345 - .000096 = .575$$

$$\text{SO POWER} = 1 - \beta = 1 - .575 = .425$$

IF THE MEAN IS ACTUALLY $\mu = 1.01$, THEN THIS TEST WILL CORRECTLY
REJECT H_0 ONLY 42.5% OF THE TIME.



POWER IS USED TO SELECT
SAMPLE SIZE NEEDED TO DETECT
A GIVEN DIFFERENCE.
TYPICALLY, INCREASE SAMPLE SIZE
TO INCREASE POWER.

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THE FOUR STEPS TO HYPOTHESIS TESTING

1. STATE HYPOTHESIS, BOTH IN WORDS (IN TERMS OF THE ALTERNATIVE)

AND IN THE FOLLOWING NOTATION:

$$H_0: \mu = \mu_0$$

$$H_a: \mu \neq \mu_0 \text{ OR } \mu < \mu_0 \text{ OR } \mu > \mu_0.$$

2. CALCULATE TEST STATISTIC

NORMAL POPULATION W/
KNOWN VARIANCE
OR LARGE SAMPLE ($n > 40$)

$$Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}$$

SMALL SAMPLE,
UNKNOWN VARIANCE
ROUGHLY NORMAL

$$T = \frac{\bar{X} - \mu_0}{S/\sqrt{n}}$$

POPULATION PROPORTION
LARGE SAMPLE ($np_0 > 5, n(1-p_0) > 5$)

$$Z = \frac{\bar{X} - np_0}{\sqrt{np_0(1-p_0)}} = \frac{\hat{P} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}}$$

3. FIND P-VALUE ASSOCIATED WITH TEST STATISTIC IN THE DIRECTION OF H_a .

[SEE P-VALUE ABOVE]

4. COMPARE P-VALUE TO α -LEVEL. IF P-VALUE $< \alpha$, REJECT H_0 .

IF P-VALUE $\geq \alpha$, FAIL TO REJECT H_0 .

STATE CONCLUSION IN TERMS OF THE ORIGINAL PROBLEM.

WHEN POSSIBLE

5. CHECK ASSUMPTIONS OF TEST: NORMALITY = ROUGHLY SYMMETRIC, UNIMODAL, ^{NO} OUTLIERS

HANDOUT 10]

EX 2: LARGE SAMPLE Z $H_a: \mu \neq 3500$ EX 3: LARGE SAMPLE Z $H_a: \mu > 0$ EX 4: LARGE SAMPLE Z $H_a: \mu < 1000$

9.3.5 1/2 Our THREE Hypothesis Tests.

Z. TEST OF POPULATION MEAN,

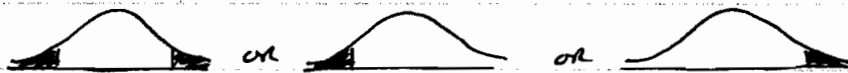
NORMAL POPULATION w/ KNOWN VARIANCE OR
LARGE SAMPLE ($n \geq 40$)

1. $H_0: \mu = \mu_0$

$H_a: \mu \neq \mu_0$ or $\mu < \mu_0$ or $\mu > \mu_0$

2. $Z = \frac{\bar{X} - \mu_0}{S/\sqrt{n}}$

3. P-VAL From TABLE II



4. CONCLUSION

T. TEST FOR POPULATION MEAN,

($n < 40$)
SMALL SAMPLE w/ UNKNOWN VARIANCE, ROUGHLY NORMAL

1. SAME AS (**Z**)

2. $T = \frac{\bar{X} - \mu_0}{S/\sqrt{n}}$

3. $df = n - 1$,
BOUNDS
P-VAL From TABLE IV

FIND VALUES IN TABLE ON df Row BORDERING T VALUE.

P-VALUE IS BETWEEN PROBABILITY BOUNDS AT TOP OF TABLE.

(FOR TWO TAILED TEST $H_a: \mu \neq \mu_0$, PROBABILITY BOUNDS ARE DOUBLED).

4. CONCLUSION.

2/2 P. TEST OF POPULATION PROPORTION

LARGE POPULATION, SAMPLE LARGE ENOUGH SO $n p_0 > 5$ AND $n(1-p_0) > 5$.

1. $H_0: p = p_0$

$H_a: p \neq p_0$ OR $p < p_0$ OR $p > p_0$

2. $Z = \frac{x - n p_0}{\sqrt{n p_0 (1-p_0)}} = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}}$, $\hat{p} = \frac{x}{n}$

3. SAME AS (Z)

4. CONCLUSION

[HANDOUT 11] EX1: T $H_a: \mu \neq 224$

EX2: T $H_a: \mu > 0$

EX3: P $H_a: p \neq .73$