

STAT 345 ELEMENTS OF MATH. STATS AND PROB THEORY
UNM

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Exam 3				

4.2 $\frac{1}{2}$ 4.2 CONTINUOUS RV'S AND PROBABILITY DISTRIBUTIONS

(4.1) Recall that a continuous RV can take on an interval of numbers.

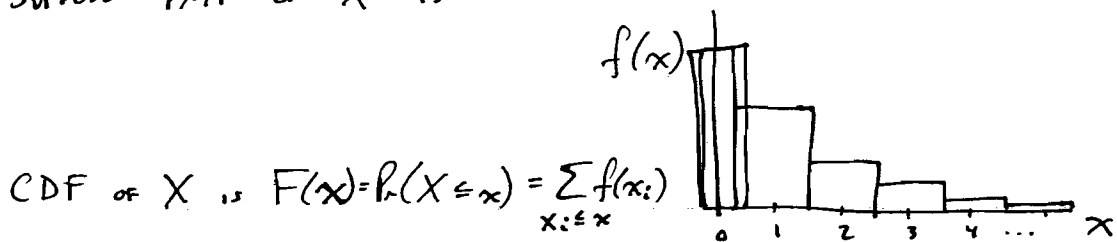
EX MEASURE THE LIFETIME OF AN ELECTRONIC COMPONENT,

$X =$ LIFETIME IN HOURS

THE RANGE OF X IS $\{x: 0 \leq x < \infty\}$

CONSIDER A DISCRETE RV X WITH RANGE $\{0, 1, 2, \dots\}$.

SUPPOSE PMF OF X IS



COMPARE AREAS OF EACH RECTANGLE WITH ~~THE~~ THE ONE YOU WANT.

AREAS SUM TO GIVE $P_r(X \leq x) = F(x)$.

TAKE RANGE TO BE $\{0, \frac{1}{5}, \frac{2}{5}, \frac{3}{5}, \frac{4}{5}, \dots\}$

CONTINUE TO REDUCE WIDTH OF RECTANGLES.

INFINITE SUM OF RECTANGLES = INTEGRAL.

TAKE SUMS AND SMALL VALUES IN ^{THE} RANGE SPACE,

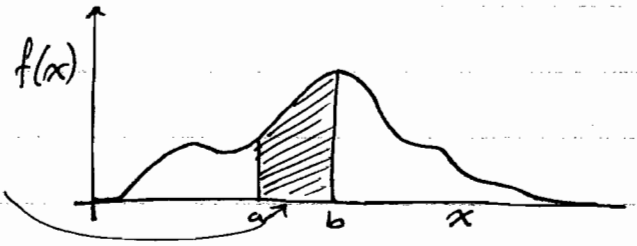
$$\sum_{x_i \leq x} f(x_i) = \int_0^x f(u) du.$$

FOR CONTINUOUS RV, THE HISTOGRAM APPROXIMATES $f(x)$.

$f(x)$ IS CALLED A PROBABILITY DENSITY FUNCTION (pdf) FOR CONTINUOUS
(PMF FOR DISCRETE)

2/2

PDF OF A CONT. RV.


 $P_r(a \leq x \leq b) = \text{Area under curve}$
PROPERTIES OF THE PDF, $f(x)$:

1. $f(x) \geq 0$
2. $\int_{-\infty}^{\infty} f(x) dx = 1$
3. $P_r(a \leq x \leq b) = \int_a^b f(x) dx = \text{Area under } f(x) \text{ from } a \text{ to } b.$

EX $P_r(x=68) = 0$

IF \uparrow ROUNDED VALUE OF A CONTINUOUS R.V.,

$P_r(67.5 \leq x < 68.5)$ MAKES MORE SENSE

NOTE: $P_r(a \leq x \leq b) = P_r(a < x \leq b) = P_r(a \leq x < b) = P_r(a < x < b) = \int_a^b f(x) dx$

EX

SUPPOSE X IS THE LIFETIME OF AN ELECTRICAL COMPONENT IN YEARS AND

THE PDF OF X IS $f(x) = \frac{1}{4} e^{-\frac{1}{4}x}$, $x \geq 0$.

- a) VERIFY $f(x)$ IS A PDF.
- b) FIND PROBABILITY THAT A RANDOMLY SELECTED COMPONENT LASTS LONGER THAN 3 YRS.
- c) FIND PROBABILITY IT LASTS BETWEEN 1 AND 2 YEARS.

a) $\int_{-\infty}^{\infty} f(x) dx = \int_0^{\infty} \frac{1}{4} e^{-\frac{1}{4}x} dx = \frac{1}{4} \int_0^{\infty} e^{-\frac{1}{4}x} dx$ $u = -\frac{1}{4}x$
 $du = -\frac{1}{4} dx \Rightarrow dx = -4 du$

$$= \frac{1}{4} \int e^u (-4 du) = - \int e^u du = -e^u = -e^{-\frac{1}{4}x} \Big|_0^{\infty}$$

$$= \lim_{x \rightarrow \infty} -e^{-\frac{1}{4}x} - (-e^0) = 0 + 1 = 1 \quad \checkmark$$

b) $P_r(X > 3) = \int_3^{\infty} \frac{1}{4} e^{-\frac{1}{4}x} dx = -e^{-\frac{1}{4}x} \Big|_3^{\infty} = 0 - (-e^{-\frac{3}{4}}) = .4724$

OR $= 1 - P_r(X \leq 3) = 1 - \int_0^3 \frac{1}{4} e^{-\frac{1}{4}x} dx$

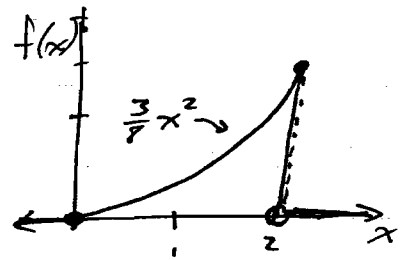
c) $P_r(1 \leq x \leq 2) = \int_1^2 \frac{1}{4} e^{-\frac{1}{4}x} dx = -e^{-\frac{1}{4}x} \Big|_1^2 = -e^{-\frac{2}{4}} - (-e^{-\frac{1}{4}}) = .1723$

4.3 CDFs.

4.3 3/1

EX THE PDF FOR A R.V. X IS GIVEN BY

$$f(x) = cx^2, \quad 0 \leq x \leq 2$$



a) Find c so that $f(x)$ is a PDF

b) Find $P_r(X \leq 3)$

c) Find the CDF of $f(x)$

$$a) \int_0^2 c \frac{3}{8} x^2 dx = 1 = c \frac{x^3}{3} \Big|_0^2 = \frac{c}{3} (2^3 - 0) = \frac{8}{3} c = 1$$

$$\Rightarrow c = \frac{3}{8}$$

b) $P_r(X \leq 3) = 1$, Acc!

c) $F(x) = P_r(X \leq x)$ CDFs ARE DEFINED ON $(-\infty, \infty)$.

$$F(x) = \begin{cases} 0, & x < 0 \\ \int_0^x \frac{3}{8} t^2 dt, & 0 \leq x \leq 2 \\ 1, & x > 2 \end{cases} \quad (\text{REPLACE } x \text{ W/ } t \text{ SO NOT SO MANY } x^2)$$

$$\int_0^x \frac{3}{8} t^2 dt = \frac{3}{8} \frac{t^3}{3} \Big|_0^x = \frac{x^3}{8}$$

$$F(x) = \begin{cases} 0, & x < 0 \\ x^3/8, & 0 \leq x \leq 2 \\ 1, & x > 2 \end{cases}$$

EX SUPPOSE R.V. X HAS PDF $f(x) = 6(x-x^2)$, $0 \leq x \leq 1$

a) Show $\int_0^1 f(x) dx = 1$.

b) Find $P_r(X > \frac{1}{2})$

c) Find cdf.

$$a) \int_0^1 6(x-x^2) dx = 6 \left(\frac{x^2}{2} - \frac{x^3}{3} \right) \Big|_0^1 = 6 \left(\frac{1}{2} - \frac{1}{3} - (0-0) \right) = 6 \left(\frac{1}{6} \right) = 1$$

$$b) P_r(X > \frac{1}{2}) = \int_{\frac{1}{2}}^1 6(x-x^2) dx = 6 \left(\frac{x^2}{2} - \frac{x^3}{3} \right) \Big|_{\frac{1}{2}}^1 = 6 \left(\frac{1}{2} - \frac{1}{3} - \left(\frac{1}{8} - \frac{1}{24} \right) \right)$$

$$= \frac{1}{2} \cdot 6 \left(\frac{1}{6} - \frac{1}{12} \right) = \frac{1}{2}$$

$$c) F(x) = P_r(X \leq x) = \begin{cases} 0, & x < 0 \\ \int_0^x 6(t-t^2) dt, & 0 \leq x \leq 1 \\ 1, & x > 1 \end{cases} \rightarrow \frac{6t^2}{2} - \frac{6t^3}{3} \Big|_0^x = 3x^2 - 2x^3$$

4.4 1/3

4.4 Mean and Variances of Continuous RV.

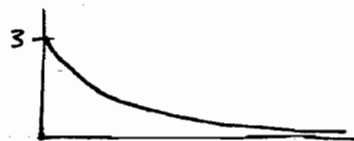
$$E(X) = \int_{-\infty}^{\infty} x f(x) dx, \quad E(g(x)) = \int_{-\infty}^{\infty} g(x) f(x) dx, \quad \text{Var}(X) = E(X^2) - (E(X))^2$$

EX Suppose X has pdf $f(x) = k e^{-3x}, x \geq 0.$ $= \int_{-\infty}^{\infty} (x-u)^2 f(x) dx$

a) Find k b) Find $F(x)$ c) Find $E(X)$ and $\text{Var}(X)$.

$$\begin{aligned} \text{a) } \int_0^{\infty} k e^{-3x} dx &= 1 = k \int_0^{\infty} e^u du = -\frac{k}{3} e^u = -\frac{k}{3} e^{-3x} \Big|_0^{\infty} \\ &= -\frac{k}{3} (0 - 1) = \frac{k}{3} = 1 \\ &\Rightarrow k = 3. \end{aligned}$$

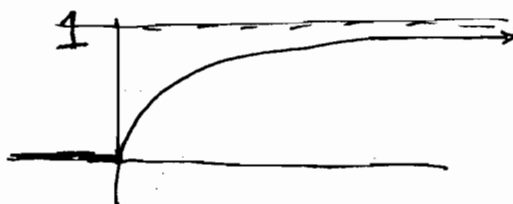
$$\text{b) } f(x) = 3e^{-3x}, x \geq 0.$$



$$F(x) = \begin{cases} 0 & x < 0 \\ \int_0^x 3e^{-3t} dt & x \geq 0 \end{cases}$$

$$\int_0^x 3e^{-3t} dt = \frac{3}{-3} e^{-3t} \Big|_0^x = -1(e^{-3x} - 1) = 1 - e^{-3x}$$

$$F(x) = \begin{cases} 0 & x < 0 \\ 1 - e^{-3x} & x \geq 0 \end{cases}$$



$$\text{c) } E(X) = \int_{-\infty}^{\infty} x f(x) dx, \quad \text{Var}(X) = E(X^2) - (E(X))^2$$

$$E(X) = \int_0^{\infty} x 3e^{-3x} dx$$

BY PARTS: ~~$\int u dv = uv - \int v du$~~ * NEXT PAGE \rightarrow

$$u = 3x; \quad dv = e^{-3x} dx$$

$$du = 3 dx; \quad v = -\frac{1}{3} e^{-3x}$$

$$E(X) = -\frac{3x}{3} e^{-3x} -$$

2/3

$$c) E(X) = \int_{-\infty}^{\infty} x f(x) dx$$

$$VAR(X) = E(X^2) - (E(X))^2$$

$$E(X) = \int_0^{\infty} 3x e^{-3x} dx = \left. \frac{3x}{-3} e^{-3x} \right|_0^{\infty} - \left. \frac{3}{9} e^{-3x} \right|_0^{\infty}$$

BY PARTS

$$\begin{array}{r} u \\ \hline 3x \end{array} \quad \begin{array}{r} dv \\ \hline e^{-3x} \end{array}$$

$$\begin{array}{r} 3 \\ \hline 0 \end{array} \quad \begin{array}{r} + \\ - \\ + \end{array} \quad \begin{array}{r} -\frac{1}{3} e^{-3x} \\ -\frac{1}{9} e^{-3x} \\ +\frac{1}{9} e^{-3x} \end{array}$$

WORKS IF ENDS
WITH ZERO

$$= (0 - 0) - \frac{1}{3} (0 - 1)$$

$$= \frac{1}{3}$$

$$E(X) = \frac{1}{3} = \mu_{MEAN}$$

$$E(X^2) = \int_0^{\infty} 3x^2 e^{-3x} dx = \left. \frac{3x^2}{-3} e^{-3x} \right|_0^{\infty} - \left. \frac{6x}{9} e^{-3x} \right|_0^{\infty} + \left. \frac{6}{27} e^{-3x} \right|_0^{\infty}$$

$$\begin{array}{r} u \\ \hline 3x^2 \end{array} \quad \begin{array}{r} dv \\ \hline e^{-3x} \end{array}$$

$$\begin{array}{r} 6x \\ \hline 6 \\ \hline 0 \end{array} \quad \begin{array}{r} + \\ - \\ + \end{array} \quad \begin{array}{r} -\frac{1}{3} e^{-3x} \\ \frac{1}{9} e^{-3x} \\ -\frac{1}{27} e^{-3x} \end{array}$$

$$= (0 - 0) - (0 - 0) + \frac{6}{27} (0 - 1)$$

$$= \frac{6}{27} = \frac{2}{9}$$

$$E(X^2) = \frac{2}{9}$$

$$VAR(X) = E(X^2) - (E(X))^2 = \frac{2}{9} - \left(\frac{1}{3}\right)^2 = \frac{2}{9} - \frac{1}{9} = \frac{1}{9}$$

$$VAR(X) = \frac{1}{9}$$

PROPERTIES OF $E(X)$ AND $VAR(X)$

IF c IS A CONSTANT,

$$E(cX) = c E(X)$$

$$E(X+c) = E(X) + c$$

$$VAR(cX) = c^2 VAR(X)$$

$$VAR(X+c) = VAR(X)$$

3/3

EX SUPPOSE THAT THE CDF OF A RV X IS GIVEN BY

$$F(x) = \begin{cases} 0, & x < 0 \\ x^3, & 0 \leq x \leq 1 \\ 1, & 1 < x \end{cases}$$

a) Find pdf $f(x)$ b) Find $E(e^x)$

$$\begin{aligned} \text{a) } f(x) &= \frac{d}{dx} F(x) = \begin{cases} \frac{d}{dx} 0, & x < 0 \\ \frac{d}{dx} x^3, & 0 \leq x \leq 1 \\ \frac{d}{dx} 1, & 1 < x \end{cases} \\ &= \begin{cases} 0, & x < 0 \\ 3x^2, & 0 \leq x \leq 1 \\ 0, & 1 < x \end{cases} \end{aligned}$$

$$f(x) = 3x^2, \quad 0 \leq x \leq 1.$$

$$\begin{aligned} \text{b) } E(e^x) &= \int_0^1 e^x (3x^2) dx = 3x^2 e^x \Big|_0^1 - 6x e^x \Big|_0^1 + 6e^x \Big|_0^1 \\ &= (3e^1 - 0) - 6(e^1 - 0) + 6(e^1 - 0) \\ &= 3e^1 - 6 \doteq 2.155 \end{aligned}$$

BY PARTS

u	dv
$3x^2$	e^x
$6x$	e^x
6	e^x
0	e^x

4.5 1/2

4.5 CONTINUOUS UNIFORM DISTRIBUTION

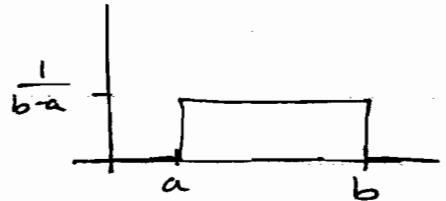
X HAS A CONTINUOUS UNIFORM DISTRIBUTION IF EVERY VALUE IN AN INTERVAL HAS THE SAME ~~PROB~~ LIKELIHOOD.

$$X \sim U(a, b)$$

$$f(x) = \frac{1}{b-a}, \quad a \leq x \leq b$$

$$E(X) = \frac{a+b}{2}$$

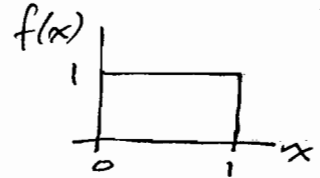
$$\text{Var}(X) = \frac{(b-a)^2}{12}$$



EX SUPPOSE YOU WANT TO RANDOMLY SELECT A NUMBER REPRESENTED BY A POINT IN THE INTERVAL $0 \leq x \leq 1$.

$X =$ NUMBER BETWEEN 0 AND 1

$$X \sim U(0, 1), \quad f(x) = 1, \quad 0 \leq x \leq 1.$$



USED TO MODEL THINGS LIKE:

- EQUALLY LIKELY WAITING TIMES
- BEAM LOADINGS
- DART BOARD EXAMPLE

EX THE THICKNESS OF STEEL SHEETS FROM ROLLING MACHINES

HAVE A UNIFORM DISTRIBUTION WITH VALUES BETWEEN 150mm - 200mm.

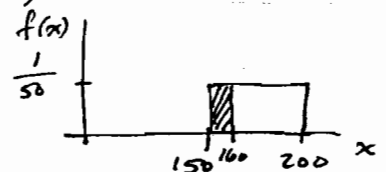
ANY SHEETS WITH THICKNESS LESS THAN 160mm ARE SCRAPPED.

a) CALCULATE THE FRACTION SCRAPPED.

b) FIND THE MEAN AND STD DEV OF THICKNESS.

$X =$ THICKNESS IN MM OF STEEL SHEET

$$X \sim U(150, 200), \quad f(x) = \frac{1}{200-150} = \frac{1}{50}, \quad 150 \leq x \leq 200$$

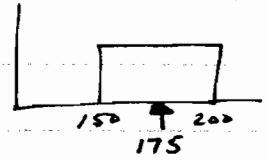


2/2

$$a) P_r(X < 160) = \int_{150}^{160} \frac{1}{50} dx = \frac{1}{50} x \Big|_{150}^{160} = \frac{1}{50} (160 - 150) = \frac{10}{50} = 0.2$$

= 20%

b) $\mu = 175$ BRUNNEN-POINT.



$$E(X) = \int_{150}^{200} x \frac{1}{50} dx = \frac{x^2}{2(50)} \Big|_{150}^{200} = \frac{200^2 - 150^2}{2(200 - 150)} = \frac{200 + 150}{2}$$
$$= \frac{b^2 - a^2}{2(b-a)} = \frac{(b-a)(b+a)}{2(b-a)} = \frac{b+a}{2} \checkmark = 175$$

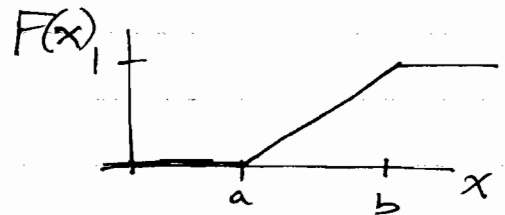
$$E(X^2) = \int_{150}^{200} x^2 \frac{1}{50} dx = \frac{1}{50} \frac{x^3}{3} \Big|_{150}^{200} = \frac{1}{150} (200^3 - 150^3) = 30833 \frac{1}{3}$$

$$VAR(X) = E(X^2) - (E(X))^2 = 30833 \frac{1}{3} - 175^2 = 208 \frac{1}{3}$$

$$STD(X) = \sqrt{\sigma^2} = \sqrt{208 \frac{1}{3}} = 14.43 \text{ mm}$$

CDF

$$F(x) = \int_{-\infty}^x f(x) dx = \begin{cases} 0, & x < a \\ \int_a^x \frac{1}{b-a} dt, & a \leq x \leq b \\ 1, & x > b \end{cases}$$
$$= \begin{cases} 0, & x < a \\ \frac{x-a}{b-a}, & a \leq x \leq b \\ 1, & x > b \end{cases}$$



4.9 1/1 4.9 EXPONENTIAL DISTRIBUTION

Recall Poisson RV counts number of successes over an interval of time or area or, etc.

For X has an exponential distribution if it measures the distance between successive counts in a Poisson process.

$$X \sim \text{Exp}(\lambda)$$

$$f(x) = \lambda e^{-\lambda x}, \quad 0 \leq x < \infty, \lambda > 0$$

$$E(X) = \frac{1}{\lambda}$$

$$\text{Var}(X) = \frac{1}{\lambda^2}$$

$$F(x) = \begin{cases} 0, & x < 0 \\ 1 - e^{-\lambda x}, & x \geq 0 \end{cases}$$

EX Suppose the number of arrivals at a service counter follows a Poisson distribution with mean number of arrivals $\lambda = 3$ per hour. Then the length of time between 2 successive arrivals, X , has $X \sim \text{Exp}(3)$, $f(x) = 3e^{-3x}$, $x \geq 0$.

$$E(X) = \frac{1}{3}, \quad \text{Var}(X) = \frac{1}{9}, \quad \sigma = \sqrt{\frac{1}{9}} = \frac{1}{3}.$$

(1st ex in 4.4)

EX The time until failure, in hours, of a fan in a computer can be modelled as $\text{Exp}(0.0003)$.

- What proportion of fans will last at least 10,000 hours?
- What proportion of fans will last at most 7,000 hours?
- What is a fan's expected lifetime?

$$X = \text{TIME UNTIL FAILURE IN HOURS}$$

$$X \sim \text{Exp}(0.0003)$$

$$a) P_r(X \geq 10,000) = \int_{10,000}^{\infty} 0.0003 e^{-0.0003x} dx = -e^{-0.0003x} \Big|_{10,000}^{\infty}$$

$$= 0 - (-e^{-3}) = e^{-3} = .0498 \approx .05$$

$$b) P_r(X \leq 7000) = F(7000) = 1 - e^{-0.0003(7000)} = 1 - e^{-2.1} = .8775 \approx .88$$

$$c) E(X) = \frac{1}{0.0003} = 3333 \frac{1}{3} \text{ hours}$$

4.6 1/3 4.6 Normal Distribution

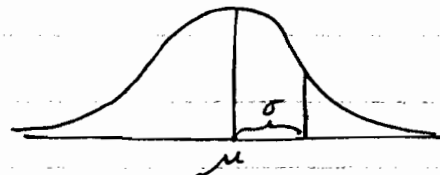
X HAS A NORMAL DISTRIBUTION IF IT ~~IS~~ ^{MEASURES} A PROCESS THAT HAS A CENTER ~~AND~~ ^{AND} IS SUBJECT TO ERROR OR SYMMETRICAL DEVIATIONS, FOR EXAMPLE. ARISES IN MANY OTHER SITUATIONS.

$$X \sim N\left(\mu, \frac{\sigma^2}{n}\right) \quad \text{where } -\infty < \mu < \infty, \sigma^2 \geq 0.$$

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(x-\mu)^2}, \quad -\infty < x < \infty$$

$$E(X) = \mu$$

$$\text{Var}(X) = \sigma^2$$



NOTE: $Z = \frac{X-\mu}{\sigma} \sim N(0, 1)$ STANDARD NORMAL.

$\Phi(z) = P(Z \leq z) =$ LOOK UP IN TABLE II, PP. 653-4.

EX $Z \sim N(0, 1)$ MEAN = $\mu = 0$, VARIANCE = $\sigma^2 = 1$.

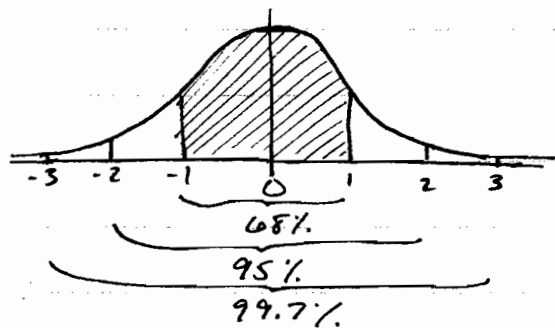
$$f(z) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(z-\mu)^2} = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}, \quad -\infty < z < \infty$$

68-95-99.7 Rule

$$P(-1 < Z < 1) = .6827 \quad (1 \text{ STD})$$

$$P(-2 < Z < 2) = .9545 \quad (2 \text{ STD})$$

$$P(-3 < Z < 3) = .9973 \quad (3 \text{ STD})$$



EX IF Z HAS A STANDARD NORMAL DISTRIBUTION, FND

a) $P(Z < 1)$

b) $P(Z < -2)$

c) $P(Z > 2)$

d) $P(Z < 0)$

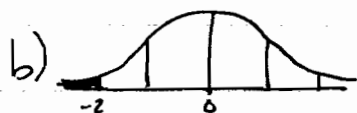


Know .6827
 \Rightarrow TAILS = $1 - .6827 = .3173$

BY SYMMETRY, $P(Z < 1) = .6827 + \frac{1}{2}(.3173)$

$$= .6827 + .1586$$

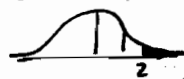
$$= .84$$



$$P(-2 < Z < 2) = .9545$$

$$P(Z < -2) = \frac{1 - .9545}{2} = .0228$$

c) SAME AS (b)



d) = $\frac{1}{2}$.

2/3

USWB TABLE II pp. 653-4.

- FIND
- $P_r(Z \leq 0.48)$
 - $P_r(Z > 1.2)$
 - $P_r(Z < -1.65)$
 - $P_r(-1.65 \leq Z \leq 0.48)$

MOTIVATION:

$$\star P_r(Z \leq 3) = \int_{-\infty}^3 f(t) dt$$


$$= \int_{-\infty}^3 \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}t^2} dt$$

NO ANTIDERIVATIVE!


LET THE COMPUTER DO IT,

⇒ USE TABLES. [REFER HANDOUT 5] COVERED AT END OF SECTION.

a) = 0.684386




b) = 1 - 0.88493 = .1151



c) = 0.049471



d) = $P_r(Z \leq 0.48) - P_r(Z < -1.65)$



= 0.684386 - 0.049471

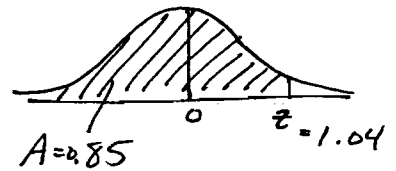
= 0.634915

EX IF $Z \sim N(0,1)$, FIND Z SUCH THAT $P_r(Z \leq z) = .85$

LOOK IN BODY OF TABLE FOR AREA = .85

CHOOSE CLOSEST Z-SCORE (VALUE OF Z).

$P_r(Z \leq 1.04) = 0.850830$

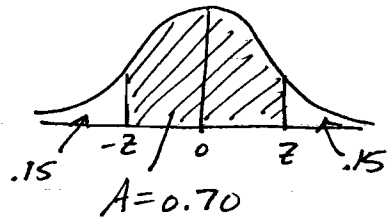


EX FIND Z SUCH THAT $P_r(-z \leq Z \leq z) = 0.70$.

CONVERT TO AN "AREA TO THE LEFT"

$P_r(Z \leq z) = .7 + .15 = .85$

(SAME AS ABOVE)

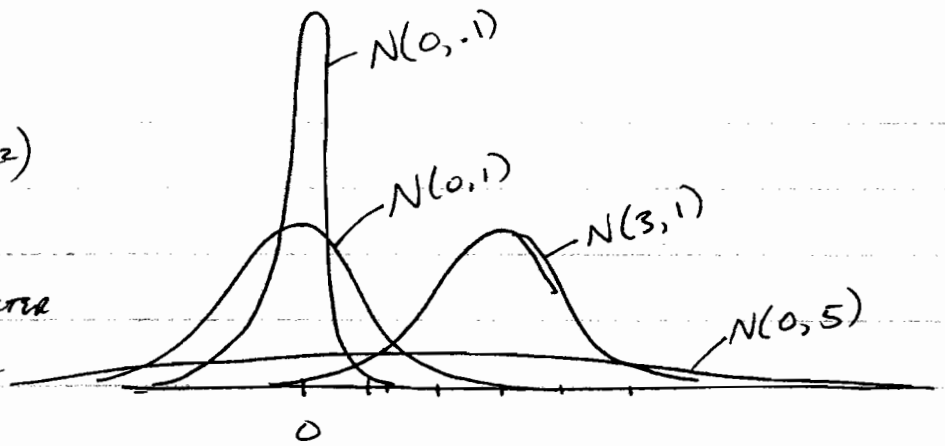


3/3

$$X \sim N(\mu, \sigma^2)$$

μ = Location Parameter

σ^2 = Scale Parameter



$$Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$$

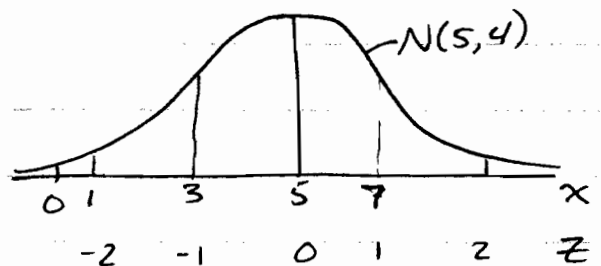
EX Suppose $X \sim N(5, 4)$, $\mu = 5$, $\sigma^2 = 4$ so $\sigma = 2$

FIND

a) ~~P~~ $P_r(X \leq 7)$

b) ~~P~~ $P_r(3 \leq X \leq 7)$

c) ~~P~~ $P_r(X \leq 7.3)$

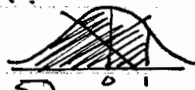


CONVERT TO Z AND USE TABLE.

a) $P_r(X \leq 7)$

$$= P_r\left(\frac{X - \mu}{\sigma} \leq \frac{7 - 5}{2}\right)$$

$$= P_r(Z \leq 1) = 0.8413$$



b) $P_r(3 \leq X \leq 7) = P_r\left(\frac{3 - 5}{2} \leq \frac{X - \mu}{\sigma} \leq \frac{7 - 5}{2}\right) = P_r(-1 \leq Z \leq 1) = .68$

c) $P_r(X \leq 7.3) = P_r\left(\frac{X - \mu}{\sigma} \leq \frac{7.3 - 5}{2}\right) = P_r(Z \leq 1.15) = 0.8749$

Handout 4]

Handout 5]

4.7 1/1

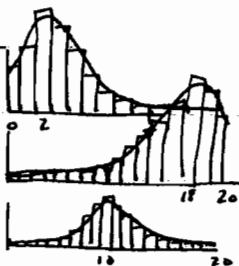
4.7 Normal Approximation to the Binomial and Poisson Distributions

Binomial Approximation

$\text{Bin}(20, .1)$

$\text{Bin}(20, .9)$

$\text{Bin}(20, .5)$



} NOT NORMAL

- LOOKS NORMAL,

$X \sim \text{Bin}(20, .5)$

CAN BE APPROXIMATED BY A NORMAL DISTRIBUTION.

$X =$ NUMBER OF SUCCESSSES OUT OF $n = 20$ TRIALS, $p = .5$

$\mu = E(X) = np$

$\sigma^2 = \text{Var}(X) = np(1-p)$

$Z = \frac{X - \mu}{\sigma} = \frac{X - np}{\sqrt{np(1-p)}} = \frac{X - 20(.5)}{\sqrt{20(.5)(1-.5)}} \sim N(0,1)$

NOW CAN USE NORMAL DIST.

PROBABILITIES FOR BINOMIAL DIST.

EX IF $X \sim \text{Bin}(20, .5)$, ESTIMATE $P_r(X \geq 3)$ USING NORMAL DISTRIBUTION.

$P_r(X \geq 3) = P_r\left(\frac{X - \mu}{\sigma} \geq \frac{3 - 20(.5)}{\sqrt{20(.5)(.5)}}\right) \approx P_r\left(Z \geq \frac{3 - 10}{\sqrt{5}}\right) = P_r(Z \geq -3.13)$
 $= .99913$



THE APPROXIMATION IS GOOD FOR IF

BOTH $np > 5$ AND $n(1-p) > 5$.

Poisson Approximation

SAME AS BINOMIAL, BUT EASIER.

NOTE $X \sim \text{Pois}(10)$

$\mu = \lambda, \sigma^2 = \lambda$



IF $X \sim \text{Pois}(\lambda)$, THEN $Z = \frac{X - \lambda}{\sqrt{\lambda}} \sim N(0,1)$ APPROXIMATELY.

GOOD WHEN $\lambda > 5$.

FOR BOTH APPROXIMATIONS, MUCH EASIER TO GET PROBABILITIES FROM TABLE II THAN ~~GETTING~~ TAKING (OFTEN LARGE) SUMMATIONS